# Equation-of-state of Neutron Stars from Gravitational Waves and Astrophysical Observations and

# How to Build a Self-Consistent Hierarchical Inference

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## **Topics**

#### Hierarchical Bayesian Inference

- What is hierarchical Bayesian inference?
- How is it connected to Directed (Acyclic) Graphs?
- How to account for selection effects?

## Neutron Star Equation of State

- Problems with phenomenological parametric EoS models
- How to
	- build (flexible) EoS models
	- condition flexible models on physics
	- discover (i.e., deal with) exotic behavior
	- extract physics without parameters

## References

#### Hierarchical Inference

- [Essick & Fishbach, ApJ \(2024\)](https://inspirehep.net/literature/2705458)
- **[Essick, PRD \(2023\)](https://inspirehep.net/literature/2674830)**
- [Essick & Holz, arXiv \(2024\)](https://inspirehep.net/literature/2808266)
- [Essick & Farr, arXiv \(2022\)](https://inspirehep.net/literature/2061444)

#### Neutron Star Equation of State

- [Landry & Essick, PRD \(2019\)](https://inspirehep.net/literature/1705987)
- [Essick, Landry, & Holz, PRD \(2020\)](https://inspirehep.net/literature/1760218)
- [Landry, Essick, & Chatziioannou \(2020\)](https://inspirehep.net/literature/1784814)
- Legred *et al* [\(including Essick\), PRD 2021\)](https://inspirehep.net/literature/1867968)
- Legred *et al* [\(including Essick\), PRD \(2022\)](https://inspirehep.net/literature/2011962)
- Essick *et al*[, PRD \(2023\)](https://inspirehep.net/literature/2659224)
- Essick *et al,* [PRC \(2020\)](https://inspirehep.net/literature/1791492)
- Essick *et al*[, PRL \(2021\)](https://inspirehep.net/literature/1847635)
- Essick *et al,* [PRC \(2021\)](https://inspirehep.net/literature/1882021)
- [Essick, ApJL \(2024\)](https://inspirehep.net/literature/2784708) *in press*

## Parametric EoS Models

## Inference of the NS EoS: systematics from parametric models **PRD** (2022)

consider a toy model:

 $\rightarrow$  fitting a 1D function (pressure vs. energy density) without constraints

linear parameterizations

point+slope  $p(\varepsilon) = p_a + c_s^2(\varepsilon - \varepsilon_a)$  $t_{\text{WOL}}$  noint

$$
p(\varepsilon) = p_a + \frac{p_b - p_a}{\varepsilon_b - \varepsilon_a} (\varepsilon - \varepsilon_a)
$$

Gaussian Process  $\vec{p} \sim \mathcal{N}(\vec{\mu}, \Sigma)$  $\Sigma_{ij} = \text{Cov}(p_i, p_j)$ 

$$
= K_{\rm se}(\varepsilon_i, \varepsilon_j) = \sigma^2 \exp \left(-\frac{(\varepsilon_i - \varepsilon_j)^2}{l^2}\right)
$$



## Inference of the NS EoS: systematics from parametric models

consider a toy model:

 $\rightarrow$  fitting a 1D function (pressure vs. energy density) without constraints

point+slope

$$
p(\varepsilon) = p_a + c_s^2(\varepsilon - \varepsilon_a)
$$

$$
p(\varepsilon) = p_a + \frac{p_b - p_a}{\varepsilon_b - \varepsilon_a} (\varepsilon - \varepsilon_a)
$$

 $GP$ 

 $\vec{p} \sim \mathcal{N}(\vec{\mu}, \Sigma)$ 

$$
\Sigma_{ij} = \text{Cov}(p_i, p_j)
$$
  
=  $K_{se}(\varepsilon_i, \varepsilon_j) = \sigma^2 \exp\left(-\frac{(\varepsilon_i - \varepsilon_j)^2}{l^2}\right)$ 



## Inference of the NS EoS: systematics from parametric models



## Non-Parametric EoS Models

## Inference of the NS EoS: nonparametric results

#### Current *Theory Agnostic* Constraints **phase** phase **phase** phase **phase** phase **phase**

transitions



## Inference of the NS EoS: nonparametric results



## *M* max  *~ 2.21* ± *0.25 M*☉ *R*(*1.4M*<sup>ⵙ</sup> ) ~ 12.5 ± 1 km

at 90% credibility









## Non-Parametric theory-informed EoS Models

## Inference of the NS EoS: incorporating low-density nuclear theory [PRC \(2020\)](https://journals.aps.org/prc/abstract/10.1103/PhysRevC.102.055803)





Inference of the NS EoS: comparing low-density theories



Inference of the NS EoS: comparing low-density theories





## Exotic TOV Sequences

#### Exotic Behavior with Efficient TOV Sequences **[ApJL \(2024\)](https://arxiv.org/pdf/2405.05395)** *in press*



## Exotic Behavior with Efficient TOV Sequences **[ApJL \(2024\)](https://arxiv.org/pdf/2405.05395)** *in press*



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## Exotic Behavior with Efficient TOV Sequences [ApJL \(2024\)](https://arxiv.org/pdf/2405.05395) in press



## Exotic Behavior with Efficient TOV Sequences **[ApJL \(2024\)](https://arxiv.org/pdf/2405.05395)** *in press*





## Extracting Physics without Parameters

Phase Transitions

















## Future Prospects: phase transitions



## Extracting Physics without Parameters

Nuclear Symmetry Energy

## Inference of the NS EoS: low-density nuclear experiment  $P_{\text{PRC (2021)}}^{\text{PRL (2021)}}$  $P_{\text{PRC (2021)}}^{\text{PRL (2021)}}$  $P_{\text{PRC (2021)}}^{\text{PRL (2021)}}$

[PRC \(2021\)](https://journals.aps.org/prc/abstract/10.1103/PhysRevC.104.065804)

Connection to "new" experimental probes: Neutron Skin Thickness ( $R_{\text{skin}}$ )

[Reed+\(2021\)](https://arxiv.org/abs/2101.03193) infer *L*  $\ge$  100 MeV based on  $R_{\text{skin}}$  = 0.29 ± 0.07 fm. Suggest this implies  $R_{1.4} \ge$  14 km.



## Inference of the NS EoS: low-density nuclear experiment

Map from nonparametric EoS in *β*-equilibrium to nuclear params describing the energy per particle near nuclear saturation ( $n_{\rm_0}$ : minimum of  $E_{_{\rm{SNM}}})$ 

 $x = n_p/n$ proton fraction solve these self-consistently  $E_{\text{nuc}}(n, x) = E_{\text{SNM}}(n) + (1 - 2x)^2 S_0(n) +$  $\mathcal{O}(x^4)$ nuclear energy per particle to obtain $(S_0(n))$  $=\frac{\varepsilon_{\beta}(n)-\varepsilon_{e}(n,x)}{n}$ and then compute  $L=3n\left(\frac{dS_0}{dn}\right)$  $E_{\text{SNM}}(n) = E_0 + \frac{1}{2} K_0 \left( \frac{n - n_0}{3n_0} \right)^2 + \cdots$ symmetric-nuclear-matter energy per particle (local min at  $n_{\text{o}}$ )  $K_{\text{sym}} = 9n^2 \left(\frac{d^2S_0}{dn^2}\right)$ condition for *β*-equilib

> constrained by astro observations (input from nonparametric analysis) measured in the lab (input from terrestrial experiment) modeled as degenerate Fermi gas (input from theory) expressed in terms of derivatives of  $E_{\text{nuc}}$

$$
\mu_i = \frac{dE}{dN_i}
$$





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astro data can distinguish between as the data can distributive control of the nuclear experiments probe lower densities nuclear theories at high densities



## Inference of the NS EoS: low-density nuclear experiment





 $L$  [MeV]

## Self-Consistent Hierarchical Bayesian Inference

## **Introduction**

Construct a hierarchical generative model that relates the differential astrophysical rate of events

$$
\frac{d\mathcal{K}}{d\theta} = \mathcal{K}p(\theta|\Lambda)
$$

to the construction of an observed catalog of discrete i.i.d. events (i.e., detected data)

$$
\{d_i,\mathbb{D}_i\}_{i=1,...,N}
$$

each of which has several latent (unobserved) parameters

 $\{\theta_i\}_{i=1,...,N}$ 

 $(\Lambda)$  population parameters (e.g., minimum and maximum mass, etc.), the

 $(\theta_i)$  single-event parameters for the i<sup>th</sup> event (e.g., masses, spins, etc.), the

 $(d_i) = n_i + h(\theta_i)$  observed data for the i<sup>th</sup> event assuming additive noise  $n_i$  and a signal model  $h(\theta_i)$ , and an

- $(\mathbb{D}_i)$  indicator signifying that the *i*<sup>th</sup> event was detected.
- $(\mathcal{K})$  expected number of astrophysical events within the past light-cone spanning the duration of the experiment or, alternatively,
- $(K) = \mathcal{K}P(\mathbb{D}|\Lambda)$  the expected number of detected events

We can represent this model as a Directed Acyclic Graph (DAG), which encodes conditional (in)dependencies between variables



## *p*(D|B,C) *p*(C|A,B) *p*(B|A) *p*(A)

or, equivalently,

We use a hierarchical model of the data-generation process



$$
p(\{\mathbb{D}_i, d_i, \theta_i\}|N, H) = \prod_i^N \underbrace{P(\mathbb{D}_i|d_i, \theta_i)\underbrace{p(d_i|\theta_i)\underbrace{p(\theta_i|H)}}_{\text{likelihood prior}}
$$

## Generative Model (bottom-up approach)

For an astronomical survey, the number of events is uncertain (Poisson distributed)

$$
p(N|K(\Lambda, \mathcal{K})) = \frac{K^N}{N!}e^{-K}
$$

## Generative Model (bottom-up approach) [ApJ \(2024\)](https://iopscience.iop.org/article/10.3847/1538-4357/ad1604) ApJ (2024)

For an astronomical survey, the number of events is uncertain (Poisson distributed)  $\mathbf{x}$ 

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and the data are i.i.d.

$$
p(\{d_i\}|\{\mathbb{D}_i\}, \Lambda, N) = \prod_i^N p(d_i|\mathbb{D}_i, \Lambda)
$$

## Generative Model (bottom-up approach) Approach and Approach Approach and Approach and Approach approach and Approach a

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p(N|K(\Lambda, \mathcal{K})) = \frac{K^N}{N!}e^{-K}
$$

and the data are i.i.d.

$$
p({d_i} | {\mathbb{D}_i}, \Lambda, N) = \prod_i^N p(d_i | {\mathbb{D}_i}, \Lambda)
$$

so that we can construct a joint distribution

 $p({d_i}, K, \Lambda | {\mathbb{D}_i}, N) \propto p(K, \Lambda) p({d_i} | {\mathbb{D}_i}, \Lambda, N) p(N|K)$ 

## Generative Model (bottom-up approach) Approach and Approach Approach and Approach and Approach approach and Approach a

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$$

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$$
p(\lbrace d_i \rbrace | \lbrace \mathbb{D}_i \rbrace, \Lambda, N) = \prod_i^N \boxed{p(d_i | \mathbb{D}_i, \Lambda)}
$$

so that we can construct a joint distribution

 $p({d_i}, K, \Lambda | {\mathbb{D}_i}, N) \propto p(K, \Lambda) p({d_i} | {\mathbb{D}_i}, \Lambda, N) p(N|K)$ 

from which we obtain the (hyper)posterior

$$
p(K, \Lambda | \{d_i, \mathbb{D}_i\}, N) = \frac{p(K, \Lambda)K^N e^{-K} \prod_i^N \left[ P(\mathbb{D}_i | \Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i | d_i, \theta) p(d_i | \theta) p(\theta | \Lambda) \right]}{\int dK d\Lambda \, p(K, \Lambda)K^N e^{-K} \prod_i^N \left[ P(\mathbb{D}_i | \Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i | d_i, \theta) p(d_i | \theta) p(\theta | \Lambda) \right]}
$$

which is *a mess…*

 $p(K,\Lambda) = p(K)p(\Lambda)$ The number of detected events is independent of the shape of the astro population  $p(K,\Lambda) = p(K)p(\Lambda)$ The number of detected events is independent of the shape of the astro population

The inference then factors

$$
p(K,\Lambda|\{d_i,\mathbb{D}_i\},N)=p(K|N)p(\Lambda|\{d_i,\mathbb{D}_i\},N)
$$

where

$$
p(K|N) = \frac{p(K)K^Ne^{-K}}{\int dK p(K)K^Ne^{-K}}
$$

$$
p(\Lambda | \{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda) \prod_i^N [P(\mathbb{D}_i | \Lambda)^{-1} \int d\theta P(\mathbb{D}_i | d_i, \theta) p(d_i | \theta) p(\theta | \Lambda)]}{\int d\Lambda p(\Lambda) \prod_i^N [P(\mathbb{D}_i | \Lambda)^{-1} \int d\theta P(\mathbb{D}_i | d_i, \theta) p(d_i | \theta) p(\theta | \Lambda)]}
$$

 $p(K,\Lambda) = p(K)p(\Lambda)$ The number of detected events is independent of the shape of the astro population

The inference then factors

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p(K,\Lambda|\{d_i,\mathbb{D}_i\},N)=p(K|N)p(\Lambda|\{d_i,\mathbb{D}_i\},N)
$$

where

$$
p(K|N) = \frac{p(K)K^{N}e^{-K}}{\int dK p(K)K^{N}e^{-K}}
$$

$$
p(\Lambda | \{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda) \prod_{i}^{N} [P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta P(\mathbb{D}_i|d_i, \theta) p(d_i|\theta) p(\theta|\Lambda)]}{\int d\Lambda p(\Lambda) \prod_{i}^{N} [P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta P(\mathbb{D}_i|d_i, \theta) p(d_i|\theta) p(\theta|\Lambda)]}
$$

We will focus on this in what follows.

Physical Detection Processes and the Inconsistency of Unphysical Model Assumptions

 $P(\mathbb{D}|d,\theta) = P(\mathbb{D}|d)$ Physical  $(\mathbb{D} \perp \theta \mid d)$ 

 $\Lambda$ 

 $\theta$ 

 $\mathbb D$ 

 $\left|d\right|$ 

$$
p(\Lambda | \{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda) \prod_i^N \left[ P(\mathbb{D}_i | \Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i | d_i) p(d_i | \theta) p(\theta | \Lambda) \right]}{\int d\Lambda p(\Lambda) \prod_i^N \left[ P(\mathbb{D}_i | \Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i | d_i) p(d_i | \theta) p(\theta | \Lambda) \right]}
$$

physical detection processes only have access to the data





 $\lfloor d \rfloor$ 



Unphysical 
$$
P(\mathbb{D}|d, \theta) = Q(\mathbb{D}|\theta)
$$
  
\n
$$
q(\Lambda | \{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda) \prod_i^N [Q(\mathbb{D}_i | \Lambda)^{-1} \int d\theta \, Q(\mathbb{D}_i | \theta) p(d_i | \theta) p(\theta | \Lambda)]}{\int d\Lambda p(\Lambda) \prod_i^N [Q(\mathbb{D}_i | \Lambda)^{-1} \int d\theta \, Q(\mathbb{D}_i | \theta) p(d_i | \theta) p(\theta | \Lambda)}
$$

incorrectly models detection as independent of the data given the event's true parameters

 $\lfloor d \rfloor$ 



$$
\begin{array}{ll}\n\text{Jnphysical} & P(\mathbb{D}|d,\theta) = Q(\mathbb{D}|\theta) & \text{no cancellation} \\
\text{on correlation} & q(\Lambda|\{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda) \prod_i^N \left[Q(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \left[Q(\mathbb{D}_i|\theta\right) p(d_i|\theta) p(\theta|\Lambda)\right]}{\int d\Lambda p(\Lambda) \prod_i^N \left[Q(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \left[Q(\mathbb{D}_i|\theta\right) p(d_i|\theta) p(\theta|\Lambda)\right]}\n\end{array}
$$

incorrectly models detection as independent of the data given the event's true parameters



$$
U1 (p1θ) = Q(ℤ|θ)
$$
\n
$$
q(Λ|{di, ℤi}, N) = \frac{p(Λ) ∏i ∑ Q(ℤi|Λ)-1 ∫ dθ Q(ℤi|θ)p(di|θ)p(di|θ)p(di|θ))}{ ∫ dΛp(Λ) ∏i ∑ Q(ℤi|Λ)-1 ∫ dθ Q(ℤi|θ)p(di|θ)p(di|θ)p(di|θ))}
$$
\n
$$
V = \frac{P(Λ) ∏i ∑ Q(ℤi|Λ)-1 ∫ dθ Q(ℤi|θ)p(di|θ)p(di|θ)p(di|θ))}
$$
\n
$$
V = \frac{P(Λ) ∩i ∑ Q(ℤ|Λ)-1 ∫ dθ Q(ℤ|θ)p(di|θ)p(di|θ)p(di|θ))}{ √ Q(ℤ|Λ)}
$$
\n
$$
q(θ|ℤ, Λ) = \frac{Q(ℤ|θ)p(θ|Λ)}{Q(ℤ|Λ)}
$$
\n
$$
V = \frac{Q(ℤ|θ)p(θ|Λ)}{Q(ℤ|Λ)}
$$
\n
$$
V = \frac{P(Λ) ∏i ∑ Q(ℤ|Λ)-1 ∫ dθ Q(ℤ|θ)p(di|θ)p(di|θ)p(di|θ))}{ √ Q(ℤ|Λ)}
$$
\n
$$
V = \frac{P(Λ) ∩i √i ∂ Q(ℤ|Λ)-1 ∫ dθ Q(ℤ|θ)p(di|θ)p(di|θ)p(di|θ))}{ √ Q(ℤ|Λ)}
$$
\n
$$
V = \frac{P(Λ) ∩i √i ∂ Q(ℤ|Λ)-1 ∫ dθ Q(ℤ|θ)p(di|θ)p(di|θ)p(di|θ))}{ √ Q(ℤ
$$

#### Common Mistake 1 Physical DAG but unphysical approximation *P*(D|Λ) ~ *Q*(D|Λ)



## Common Mistake 1 Physical DAG but unphysical approximation *P*(D|Λ) ~ *Q*(D|Λ)

truth  $p(m_1|\Lambda) = \mathcal{N}(\mu_{m_1}, \sigma_{m_1}^2) \Theta(m_{\min} \leq m_1 \leq m_{\max})$ physical  $p(m_2|m_1,\Lambda) \propto \Theta(m_{\min} \leq m_2 \leq m_1)$ unphysical  $p(z|\Lambda) \propto \frac{dV_c}{dz} (1+z)^{\kappa-1}$ 0.48  $\sigma_{m_1}~[M_\odot]$  $\mathcal{O}_{\mathcal{W}}$ selection is primarily against high-*z* systems, so the bias primarily affects the redshift-evolution  $\circ \mathbb{P}$ 0.24  $Q(\mathbb{D}|z) \ll P(\mathbb{D}|z) \Rightarrow \kappa \uparrow$  $\circ \frac{1}{6}$ in order to keep  $\varphi$  $K = \mathcal{K}P(\mathbb{D}|\Lambda) = \mathcal{K}\int d\theta\, p(\theta|\Lambda)P(\mathbb{D}|\theta)$  $\mathcal{L}$ approximately constant Z.  $\circ$  $\hat{\mathcal{L}}$  $\mathcal{S}_{\mathcal{L}}$  $\mathcal{S}^5$ 1.50 0.16 0.24 0.32 0.10 0.18  $\mathcal{L}$  $\hat{\mathcal{L}}$  $\mathcal{O}$  $\breve{\mathcal{C}}$  $\mathcal{O}_{\lambda}$ [arXiv:2310.02017 \(2023\)](https://arxiv.org/abs/2310.02017)<br> $\mu_{m_1}$  [M<sub>o</sub>]  $\sigma_{m_1}$  [M<sub> $\odot$ </sub>]  $\kappa$ 

[ApJ \(2024\)](https://iopscience.iop.org/article/10.3847/1538-4357/ad1604)

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Motivated by the observation

$$
p(\theta|\Lambda) = \left(\frac{p(\theta|\mathbb{D}, \Lambda)}{P(\mathbb{D}|\theta)}\right) P(\mathbb{D}|\Lambda)
$$

#### Common Mistake 2 Fitting the "detected distribution" and then dividing by *P*(D|θ)

Motivated by the observation

$$
p(\theta|\Lambda) = \left(\frac{p(\theta|\mathbb{D}, \Lambda)}{P(\mathbb{D}|\theta)}\right) P(\mathbb{D}|\Lambda)
$$

It is tempting to fit for the distribution of true-parameters for detected events *p*(θ|D,Λ) from the observed data via

$$
q(\lbrace d_i \rbrace | \lbrace D_i \rbrace, N, \Lambda) = \prod_i^N \int d\theta \, p(d_i | \theta) q(\theta | \mathbb{D}_i, \Lambda)
$$

with a "flexible enough" model for *q*(θ|D,Λ)

#### Common Mistake 2 Fitting the "detected distribution" and then dividing by *P*(D|θ)

Motivated by the observation

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p(\theta|\Lambda) = \left(\frac{p(\theta|\mathbb{D}, \Lambda)}{P(\mathbb{D}|\theta)}\right) P(\mathbb{D}|\Lambda)
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It is tempting to fit for the distribution of true-parameters for detected events *p*(θ|D,Λ) from the observed data via

$$
q(\{d_i\}|\{D_i\},N,\Lambda)=\prod_i^N\int d\theta\, p(d_i|\theta)q(\theta|\mathbb{D}_i,\Lambda)
$$

with a "flexible enough" model for *q*(θ|D,Λ) and then estimate the astro distribution

$$
q(\theta|\Lambda) \propto \frac{q(\theta|\mathbb{D}, \Lambda)}{P(\mathbb{D}|\theta)}
$$

However, this procedure yields *q*(θ|D,Λ) that are too narrow, and therefore biased estimates of *q*(θ|Λ).

## Common Mistake 2

Fitting the "detected distribution" and then dividing by *P*(D|θ)



$$
p(\delta\phi|\Lambda) = \mathcal{N}(\mu_{\Lambda}, \sigma_{\Lambda}^2) \qquad p(\delta\hat{\phi}|\delta\phi) = \mathcal{N}(\delta\phi, \sigma_{o}^2) \qquad P(\mathbb{D}|\delta\hat{\phi}) = \exp\left(-\frac{(\delta\hat{\phi} - \mu_{D})^2}{\sigma_{D}^2}\right)
$$

[ApJ \(2024\)](https://iopscience.iop.org/article/10.3847/1538-4357/ad1604)

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