# Equation-of-state of Neutron Stars from Gravitational Waves and Astrophysical Observations and

# How to Build a Self-Consistent Hierarchical Inference

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### Topics

#### Hierarchical Bayesian Inference

- What is hierarchical Bayesian inference?
- How is it connected to Directed (Acyclic) Graphs?
- How to account for selection effects?

### Neutron Star Equation of State

- Problems with phenomenological parametric EoS models
- How to
  - build (flexible) EoS models
  - condition flexible models on physics
  - discover (i.e., deal with) exotic behavior
  - extract physics without parameters

### References

#### **Hierarchical Inference**

- Essick & Fishbach, ApJ (2024)
- Essick, PRD (2023)
- Essick & Holz, arXiv (2024)
- Essick & Farr, arXiv (2022)

#### Neutron Star Equation of State

- Landry & Essick, PRD (2019)
- Essick, Landry, & Holz, PRD (2020)
- Landry, Essick, & Chatziioannou (2020)
- Legred *et al* (including Essick), PRD 2021)
- Legred *et al* (including Essick), PRD (2022)
- Essick et al, PRD (2023)
- Essick et al, PRC (2020)
- Essick *et al*, PRL (2021)
- Essick et al, PRC (2021)
- Essick, ApJL (2024) in press

### Parametric EoS Models

### Inference of the NS EoS: systematics from parametric models

consider a toy model:

 $\rightarrow$  fitting a 1D function (pressure vs. energy density) without constraints

linear parameterizations

point+slope  $p(\varepsilon) = p_a + c_s^2(\varepsilon - \varepsilon_a)$ two-point  $p_b - p_a$ 

$$p(\varepsilon) = p_a + \frac{p_b - p_a}{\varepsilon_b - \varepsilon_a} (\varepsilon - \varepsilon_a)$$

Gaussian Process  $\vec{p} \sim \mathcal{N}(\vec{\mu}, \Sigma)$  $\Sigma_{ij} = \text{Cov}(p_i, p_j)$ 

$$= K_{\rm se}(\varepsilon_i, \varepsilon_j) = \sigma^2 \exp\left(-\frac{(\varepsilon_i - \varepsilon_j)^2}{l^2}\right)$$



### Inference of the NS EoS: systematics from parametric models

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point+slope

$$p(\varepsilon) = p_a + c_s^2(\varepsilon - \varepsilon_a)$$

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GP

 $\vec{p} \sim \mathcal{N}(\vec{\mu}, \Sigma)$ 

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=  $K_{se}(\varepsilon_i, \varepsilon_j) = \sigma^2 \exp\left(-\frac{(\varepsilon_i - \varepsilon_j)^2}{l^2}\right)$ 



### Inference of the NS EoS: systematics from parametric models



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### Non-Parametric EoS Models

### Inference of the NS EoS: nonparametric results

#### Current Theory Agnostic Constraints

phase transitions



### Inference of the NS EoS: nonparametric results

	Observable	Prior	w/ PSRs	w/o J0740 + 6620	$w/J0740{+}6620$	
					Miller+	Riley+
Properties of the EoS Properties defined for both NSs and BHs	$M_{ m max} \; [{ m M}_{\odot}]$	$1.47\substack{+0.71 \\ -1.37}$	$2.24\substack{+0.48 \\ -0.24}$	$2.20\substack{+0.30 \\ -0.19}$	$2.21^{+0.31}_{-0.21}$	$2.19\substack{+0.27 \\ -0.19}$
	$p( ho_{ m nuc}) \; [10^{33} { m dyn/cm^2}]$	$2.25^{+5.81}_{-2.15}$	$6.07\substack{+7.53 \\ -5.93}$	$4.05\substack{+3.59 \\ -3.74}$	$4.30^{+3.37}_{-3.80}$	$4.15_{-3.76}^{+3.50}$
	$p(2\rho_{\rm nuc}) \ [10^{34} {\rm dyn/cm^2}]$	$1.22^{+4.86}_{-1.21}$	$6.00\substack{+4.79 \\ -5.99}$	$3.75^{+2.36}_{-2.98}$	$4.38^{+2.46}_{-2.96}$	$3.90^{+2.11}_{-2.88}$
	$p(6 ho_{ m nuc})~[10^{35}{ m dyn/cm^2}]$	$2.43_{-2.43}^{+4.70}$	$7.51_{-5.15}^{+6.77}$	$8.33_{-4.14}^{+5.22}$	$7.41_{-4.18}^{+5.87}$	$7.82^{+5.47}_{-3.53}$
	$\max\left\{c_s^2/c^2\right\} \   \ \rho \le \rho_c(M_{\max})$	$0.76\substack{+0.24 \\ -0.37}$	$0.72\substack{+0.28 \\ -0.26}$	$0.84^{+0.16}_{-0.28}$	$0.75\substack{+0.25 \\ -0.24}$	$0.80^{+0.20}_{-0.26}$
	$ ho\left(\max\left\{c_{s}^{2}/c^{2} ight\} ight)\left[10^{15}{ m g/cm^{3}} ight]$	$1.38^{+1.65}_{-1.34}$	$0.97\substack{+0.64 \\ -0.70}$	$1.13\substack{+0.64\\-0.63}$	$1.01\substack{+0.63 \\ -0.53}$	$1.10\substack{+0.63 \\ -0.58}$
	$p\left(\max\left\{c_{s}^{2}/c^{2} ight\} ight)$ [10 <sup>35</sup> dyn/cm <sup>2</sup> ]	$1.65\substack{+8.16 \\ -1.65}$	$2.68\substack{+5.18 \\ -2.68}$	$3.52^{+6.90}_{-3.48}$	$2.77^{+5.81}_{-2.70}$	$3.26^{+6.51}_{-3.15}$
	$R_{1.4}$ [km]	$8.09\substack{+5.68\\-3.96}$	$13.54_{-3.13}^{+2.61}$	$12.25_{-1.33}^{+1.13}$	$12.56^{+1.00}_{-1.07}$	$12.34^{+1.01}_{-1.25}$
	$R_{2.0}  \mathrm{[km]}$	$5.90\substack{+6.97 \\ -0.00}$	$13.18\substack{+3.02 \\ -2.90}$	$12.05_{-1.45}^{+1.18}$	$12.41^{+1.00}_{-1.10}$	$12.09^{+1.07}_{-1.17}$
	$\Delta R \equiv R_{2.0} - R_{1.4} \; [\rm km]$	$0.48^{+1.28}_{-6.67}$	$-0.07\substack{+1.00\\-1.04}$	$-0.17\substack{+0.85\\-0.83}$	$-0.12\substack{+0.83 \\ -0.85}$	$-0.20\substack{+0.82\\-0.88}$
	$\Lambda_{1.4}$	$24^{+841}_{-24}$	$795^{+1262}_{-708}$	$442^{+235}_{-274}$	$507^{+234}_{-242}$	$457^{+219}_{-256}$
	$\Lambda_{2.0}$	$0^{+54}_{-0}$	$66^{+184}_{-66}$	$34^{+35}_{-27}$	$44^{+34}_{-30}$	$35^{+32}_{-24}$
Properties defined only for NSs	$ ho_{ m c}(1.4{ m M}_{\odot})~[10^{14}{ m g/cm}^3]$	$8.4^{+12.5}_{-6.0}$	$5.7^{+3.2}_{-3.1}$	$7.2^{+2.6}_{-1.7}$	$6.7^{+1.7}_{-1.3}$	$7.1^{+2.1}_{-1.5}$
	$ ho_{ m c}(2.0{ m M}_{\odot})~[10^{14}{ m g/cm^3}]$	$9.0^{+5.7}_{-6.3}$	$8.5^{+4.8}_{-5.3}$	$10.5^{+4.1}_{-3.8}$	$9.7\substack{+3.6 \\ -3.1}$	$10.4^{+3.6}_{-3.5}$
	$ ho_{ m c}(M_{ m max})~[10^{15}{ m g/cm}^3]$	$2.4^{+0.9}_{-2.0}$	$1.4_{-0.6}^{+0.5}$	$1.6^{+0.3}_{-0.4}$	$1.5\substack{+0.3 \\ -0.4}$	$1.6\substack{+0.3 \\ -0.3}$

 $M_{\rm max} \sim 2.21 \pm 0.25 M_{\odot}$  $R(1.4M_{\odot}) \sim 12.5 \pm 1 \text{ km}$ 

at 90% credibility







### Non-Parametric theory-informed EoS Models

### Inference of the NS EoS: incorporating low-density nuclear theory





Inference of the NS EoS: comparing low-density theories





Inference of the NS EoS: comparing low-density theories



## Exotic TOV Sequences

ApJL (2024) in press





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### **Extracting Physics without Parameters**

**Phase Transitions** 

















### Future Prospects: phase transitions



### **Extracting Physics without Parameters**

Nuclear Symmetry Energy

### Inference of the NS EoS: low-density nuclear experiment

PRL (2021) PRC (2021)

Connection to "new" experimental probes: Neutron Skin Thickness (R<sub>ekin</sub>)

Reed+(2021) infer  $L \ge 100$  MeV based on  $R_{skin} = 0.29 \pm 0.07$  fm. Suggest this implies  $R_{14} \ge 14$  km.



### Inference of the NS EoS: low-density nuclear experiment

Map from nonparametric EoS in  $\beta$ -equilibrium to nuclear params describing the energy per particle near nuclear saturation ( $n_0$ : minimum of  $E_{SNM}$ )

 $x = n_p/n$ proton fraction solve these self-consistently  $E_{\rm nuc}(n,x) = E_{\rm SNM}(n) + (1-2x)^2 S_0(n) +$  $\mathcal{O}(x^4)$ nuclear energy per particle to obtain  $(S_0(n))$  $= \frac{\varepsilon_{\beta}(n) - \varepsilon_{e}(n, x)}{\varepsilon_{\beta}(n, x)} - \varepsilon_{e}(n, x)$ and then compute  $L = 3n\left(\frac{dS_0}{dn}\right)$  $E_{\rm SNM}(n) = E_0 + \frac{1}{2} K_0 \left( \frac{n - n_0}{3n_0} \right)^2 + \cdots$ symmetric-nuclear-matter energy per particle (local min at  $n_{0}$ )  $K_{\rm sym} = 9n^2 \left(\frac{d^2 S_0}{dn^2}\right)$ condition for  $\beta$ -equilib

> constrained by astro observations (input from nonparametric analysis) measured in the lab (input from terrestrial experiment) modeled as degenerate Fermi gas (input from theory) expressed in terms of derivatives of  $E_{nuc}$

$$\mu_i = \frac{dE}{dN_i}$$







nuclear experiments probe lower densities



### Inference of the NS EoS: low-density nuclear experiment





 $L \,[{\rm MeV}]$ 

### Self-Consistent Hierarchical Bayesian Inference

Construct a hierarchical generative model that relates the differential astrophysical rate of events

$$\frac{d\mathcal{K}}{d\theta} = \mathcal{K}p(\theta|\Lambda)$$

to the construction of an observed catalog of discrete i.i.d. events (i.e., detected data)

$$\{d_i, \mathbb{D}_i\}_{i=1,\dots,N}$$

each of which has several latent (unobserved) parameters

 $\{\theta_i\}_{i=1,\dots,N}$ 

(A) population parameters (e.g., minimum and maximum mass, etc.), the

 $(\theta_i)$  single-event parameters for the  $i^{\text{th}}$  event (e.g., masses, spins, etc.), the

 $(d_i) = n_i + h(\theta_i)$  observed data for the *i*<sup>th</sup> event assuming additive noise  $n_i$  and a signal model  $h(\theta_i)$ , and an

- $(\mathbb{D}_i)$  indicator signifying that the  $i^{\text{th}}$  event was detected.
- $(\mathcal{K})$  expected number of astrophysical events within the past light-cone spanning the duration of the experiment or, alternatively,
- $(K) = \mathcal{K}P(\mathbb{D}|\Lambda)$  the expected number of detected events

We can represent this model as a Directed Acyclic Graph (DAG), which encodes conditional (in)dependencies between variables



#### equation

### $p(\mathsf{D}|\mathsf{B},\mathsf{C}) p(\mathsf{C}|\mathsf{A},\mathsf{B}) p(\mathsf{B}|\mathsf{A}) p(\mathsf{A})$

We use a hierarchical model of the data-generation process



or, equivalently,

$$p(\{\mathbb{D}_i, d_i, \theta_i\} | N, \mathbf{H}) = \prod_{i}^{N} \underbrace{P(\mathbb{D}_i | d_i, \theta_i) p(d_i | \theta_i) p(\theta_i | \mathbf{H})}_{\text{likelihood prior}}$$

For an astronomical survey, the number of events is uncertain (Poisson distributed)

$$p(N|K(\Lambda,\mathcal{K})) = \frac{K^N}{N!}e^{-K}$$

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so that we can construct a joint distribution

 $p(\{d_i\}, K, \Lambda | \{\mathbb{D}_i\}, N) \propto p(K, \Lambda) p(\{d_i\} | \{\mathbb{D}_i\}, \Lambda, N) p(N | K)$ 

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from which we obtain the (hyper)posterior

$$p(K,\Lambda|\{d_i,\mathbb{D}_i\},N) = \frac{p(K,\Lambda)K^N e^{-K} \prod_i^N \left[ P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta P(\mathbb{D}_i|d_i,\theta) p(d_i|\theta) p(\theta|\Lambda) \right]}{\int dK d\Lambda \, p(K,\Lambda) K^N e^{-K} \prod_i^N \left[ P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i|d_i,\theta) p(d_i|\theta) p(\theta|\Lambda) \right]}$$

which is a mess ...

 $p(K,\Lambda) = p(K)p(\Lambda)$  The number of detected events is independent of the shape of the astro population

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The inference then factors

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where

$$p(K|N) = \frac{p(K)K^N e^{-K}}{\int dK \, p(K)K^N e^{-K}}$$

$$p(\Lambda|\{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda)\prod_i^N \left[P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i|d_i, \theta) p(d_i|\theta) p(\theta|\Lambda)\right]}{\int d\Lambda p(\Lambda)\prod_i^N \left[P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \, P(\mathbb{D}_i|d_i, \theta) p(d_i|\theta) p(\theta|\Lambda)\right]}$$

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We will focus on this in what follows.

Physical Detection Processes and the Inconsistency of Unphysical Model Assumptions

 $\begin{array}{ll} \text{Physical} & P(\mathbb{D}|d,\theta) = P(\mathbb{D}|d) \\ & {}_{(\mathbb{D}\,\perp\,\theta\,\mid\,d)} \end{array}$ 

 $\Lambda$ 

 $\theta$ 

 $\mathbb{D}$ 

d

$$p(\Lambda|\{d_i, \mathbb{D}_i\}, N) = \frac{p(\Lambda) \prod_i^N \left[ P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \ P(\mathbb{D}_i|d_i) p(d_i|\theta) p(\theta|\Lambda) \right]}{\int d\Lambda p(\Lambda) \prod_i^N \left[ P(\mathbb{D}_i|\Lambda)^{-1} \int d\theta \ P(\mathbb{D}_i|d_i) p(d_i|\theta) p(\theta|\Lambda) \right]}$$

physical detection processes only have access to the data





d



incorrectly models detection as independent of the data given the event's true parameters

d



$$\begin{array}{ll} \text{Unphysical} & P(\mathbb{D}|d,\theta) = Q(\mathbb{D}|\theta) & \text{no cancellation} \\ & (\mathbb{D} \perp d \mid \theta) & q(\Lambda|\{d_i,\mathbb{D}_i\},N) = \frac{p(\Lambda)\prod_i^N \left[Q(\mathbb{D}_i|\Lambda)^{-1}\int d\theta Q(\mathbb{D}_i|\theta)p(d_i|\theta)p(\theta|\Lambda)\right]}{\int d\Lambda p(\Lambda)\prod_i^N \left[Q(\mathbb{D}_i|\Lambda)^{-1}\int d\theta Q(\mathbb{D}_i|\theta)p(d_i|\theta)p(\theta|\Lambda)\right]} \end{array}$$

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#### Common Mistake 1 Physical DAG but unphysical approximation $P(D|\Lambda) \sim Q(D|\Lambda)$



### Common Mistake 1 Physical DAG but unphysical approximation $P(D|\Lambda) \sim Q(D|\Lambda)$

truth  $p(m_1|\Lambda) = \mathcal{N}(\mu_{m_1}, \sigma_{m_1}^2) \Theta(m_{\min} \le m_1 \le m_{\max})$ physical  $p(m_2|m_1,\Lambda) \propto \Theta(m_{\min} \leq m_2 \leq m_1)$ unphysical  $p(z|\Lambda) \propto \frac{dV_c}{dz}(1+z)^{\kappa-1}$ 0.10  $\sigma_{m_1} \; [M_\odot]$ 0,10 0.32 0.24 0.16 in order to keep 0 5 2 app 0 5 20 1.350 1.50 0.10 0.24 0.32 0.40 0.4° 20 5 0 5 0, arXiv:2310.02017 (2023)  $\mu_{m_1} [M_{\odot}]$  $\sigma_{m_1} [M_{\odot}]$ к

ApJ (2024)

selection is primarily against high-z systems, so the bias primarily affects the redshift-evolution

$$Q(\mathbb{D}|z) \ll P(\mathbb{D}|z) \Rightarrow \kappa \uparrow$$

$$K = \mathcal{K}P(\mathbb{D}|\Lambda) = \mathcal{K}\int d\theta\, p(\theta|\Lambda)P(\mathbb{D}|\theta)$$
 proximately constant

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Motivated by the observation

$$p(\boldsymbol{\theta}|\boldsymbol{\Lambda}) = \left(\frac{p(\boldsymbol{\theta}|\mathbb{D},\boldsymbol{\Lambda})}{P(\mathbb{D}|\boldsymbol{\theta})}\right) P(\mathbb{D}|\boldsymbol{\Lambda})$$

#### Common Mistake 2 Fitting the "detected distribution" and then dividing by $P(D|\theta)$

Motivated by the observation

$$p(\theta|\Lambda) = \left(\frac{p(\theta|\mathbb{D},\Lambda)}{P(\mathbb{D}|\theta)}\right) P(\mathbb{D}|\Lambda)$$

It is tempting to fit for the distribution of true-parameters for detected events  $p(\theta|D,\Lambda)$  from the observed data via

$$q(\{d_i\}|\{D_i\},N,\Lambda) = \prod_i^N \int d\theta \, p(d_i|\theta) q(\theta|\mathbb{D}_i,\Lambda)$$

with a "flexible enough" model for  $q(\theta|D,\Lambda)$ 

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with a "flexible enough" model for  $q(\theta|D,\Lambda)$  and then estimate the astro distribution

$$q(\boldsymbol{\theta}|\boldsymbol{\Lambda}) \propto \frac{q(\boldsymbol{\theta}|\mathbb{D},\boldsymbol{\Lambda})}{P(\mathbb{D}|\boldsymbol{\theta})}$$

However, this procedure yields  $q(\theta|D,\Lambda)$  that are too narrow, and therefore biased estimates of  $q(\theta|\Lambda)$ .

### Common Mistake 2

Fitting the "detected distribution" and then dividing by  $P(D|\theta)$ 



$$p(\delta\phi|\Lambda) = \mathcal{N}(\mu_{\Lambda}, \sigma_{\Lambda}^2) \qquad p(\delta\hat{\phi}|\delta\phi) = \mathcal{N}(\delta\phi, \sigma_o^2) \qquad P(\mathbb{D}|\delta\hat{\phi}) = \exp\left(-\frac{(\delta\hat{\phi} - \mu_D)^2}{\sigma_D^2}\right)$$

ApJ (2024)

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