Machine learning for extreme values

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- Analysis of rare phenomena with small probabilities
- Impact on various risks (health, environment, economy,...)



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- Causal inference for extremes [3] [4]
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[2] S. Engelke and J. Ivanovs. "Sparse structures for multivariate extremes". In: Annu. Rev. Stat. Appl. 8 (2021), pp. 241–270.

<sup>[1]</sup> S. Engelke and A. S. Hitz. "Graphical models for extremes (with discussion)". In: J. R. Stat. Soc. Ser. B. Stat. Methodol. 82.4 (2020), pp. 871–932.

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- Machine learning: prediction beyond the data range

- [1] S. Engelke and A. S. Hitz. "Graphical models for extremes (with discussion)". In: J. R. Stat. Soc. Ser. B. Stat. Methodol. 82.4 (2020), pp. 871–932.
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- [3] D. Deuber, J. Li, S. Engelke, and M. H. Maathuis. "Estimation and inference of extremal quantile treatment effects for heavy-tailed distributions". In: J. Amer. Statist. Assoc. (2023). to appear.
- [4] N. Gnecco, N. Meinshausen, J. Peters, and S. Engelke. "Causal discovery in heavy-tailed models". In: Ann. Statist. 49.3 (2021), pp. 1755–1778.
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– Prediction of conditional quantile at level  $\tau \in (0,1)$ :

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- If  $\tau$  is close to 1, we speak of extreme quantile regression
- Again, classical machine learning methods perform poorly

# Extreme quantile regression: Aare river in Bern



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– Late warnings during flood in August 2005



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– Late warnings during flood in August 2005



- $-$  Y is daily discharge at Bern station
- X contains discharge and precipitation from previous days, etc.

 $-$  For independent copies  $({\bf X}_1,Y_1),\ldots,({\bf X}_n,Y_n)$  of  ${\bf X}\in\mathbb{R}^p$  and  $\,\in\mathbb{R},$  the goal is to predict the conditional quantile at level  $\tau \in (0,1)$ 

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<sup>[5]</sup> V. Chernozhukov. "Extremal quantile regression". In: The Annals of Statistics 33.2 (2005), pp. 806–839. [6] V. Chavez-Demoulin and A. C. Davison. "Generalized additive modelling of sample extremes". In: Journal of the Royal Statistical Society: Series C (Applied Statistics) 54.1 (2005), pp. 207–222.

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- There are different scenarios depending on the quantile level  $\tau = \tau_n$ :
	- $\tau_n \equiv \tau_0 < 1$  (classical case)
	- $\tau_n \rightarrow 1$ , and  $n(1 \tau_n) \rightarrow \infty$  (intermediate case)
	- $-\tau_n \to 1$ , and  $n(1-\tau_n) \to c \in [0,\infty)$  (extreme case)

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- Classical methods for quantile regression only work well in the case of fixed  $\tau_n \equiv \tau_0 < 1$
- $-$  Existing methods from extreme value theory are not flexible enough or do not generalize well to higher dimensions [5] [6]
- Goal: Develop a new method for extreme quantile regression that works well with high-dimensional and complex data

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\mathbb{P}(Y>y)=\mathbb{P}(Y>u)\times\mathbb{P}(Y>y\mid Y>u)
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$$
  
 
$$
\approx \mathbb{P}(Y > u) \times (1 - H_{\sigma,\gamma}(y - u))
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where  $H_{\sigma,\gamma}$  is the cdf of the GPD with scale  $\sigma > 0$ and shape  $\gamma \in \mathbb{R}$ .



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The sign of  $\gamma$  characterizes the domain of attraction (DOA) of Y

- $-\gamma > 0$ : Y has heavy tails (Fréchet DOA)
- $-\gamma = 0$ : Y has light tails (Gumbel DOA)
- $-\gamma$  < 0: Y has finite upper endpoint (Weibull DOA)

## **Estimation**

- Consider i.i.d. data  $Y_1, \ldots, Y_n$  and estimate empirically the quantile  $u = \hat{Q}(\tau_0)$  for an intermediate quantile level  $\tau_0 < 1$ .
- Define the exceedances above the threshold as

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Estimate the parameters by maximum likelihood

$$
\hat{\theta} = \underset{\theta = (\sigma, \gamma)}{\arg \max} \sum_{i=1}^{n} \ell_{Z_i}(\theta).
$$

## Extreme quantile estimation

 $-$  Approximation of quantile  $Q(\tau)=F_Y^{-1}(\tau)$  for probability level  $\tau>\tau_0$  by inverting the cdf  $H_{\hat\sigma,\hat\gamma}$  of the GPD

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– Compute a T-year event as  $\hat{Q}(1 - 1/(n_Y T))$ , where  $n_Y$  is the number of observations per year.



## During the 2005 flood in Bern



– Daily observations during the 2005 flood in Bern together with 100-year return level estimate.

– Assume the GPD model

$$
(Y - \hat{Q}(\tau_0) | Y > \hat{Q}(\tau_0), \qquad ) \sim H_{\sigma_-, \gamma}
$$

where  $\tau_0$  is an intermediate quantile level

– Assume the conditional GPD model

$$
(Y - \hat{Q}_x(\tau_0) \mid Y > \hat{Q}_x(\tau_0), \mathbf{X} = \mathbf{x}) \sim H_{\sigma(\mathbf{x}), \gamma(\mathbf{x})}
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where  $\tau_0$  is an intermediate quantile level, and  $\hat{Q}_x(\tau_0)$  is an estimate of the conditional  $\tau_0$  quantile of  $Y \mid X = x$ 

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- $\hat{Q}_x(\tau_0)$  can be estimated with classical methods, e.g., a quantile random forest.
- For an extreme level  $\tau > \tau_0$  we can estimate

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\hat{Q}_{\mathsf{x}}(\tau) = \hat{Q}_{\mathsf{x}}(\tau_0) + \hat{\sigma}(\mathsf{x})\frac{\left(\frac{1-\tau}{1-\tau_0}\right)^{-\hat{\gamma}(\mathsf{x})}-1}{\hat{\gamma}(\mathsf{x})},
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where  $\hat{\theta}(\mathbf{x}) = (\hat{\sigma}(\mathbf{x}), \hat{\gamma}(\mathbf{x}))$  is an estimate of the conditional GPD parameters.

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– The triple  $(\hat{Q}_x(\tau_0), \hat{\sigma}(x), \hat{\gamma}(x))$  provides a model for the tail of Y |  $X = x$ .
– Generalized random forests (GRF) [7] are ensemble methods based on trees that can be seen as adaptive nearest neighbors that optimize

$$
\hat{\theta}(\mathbf{x}) = \arg\min_{\theta} \sum_{i=1}^{n} w_n(\mathbf{x}, \mathbf{X}_i) L(\theta, Y_i),
$$

where L is a loss function and  $w_n(x, X_i)$  are localizing weights from the GRF

<sup>[7]</sup> S. Athey, J. Tibshirani, and S. Wager. "Generalized random forests". In: Ann. Statist. 47.2 (2019), pp. 1148–1178.

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 $-$  Least squares regression: loss  $L(\theta,Y) = (Y - \theta)^2$  and

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– For extreme quantile levels  $\tau \approx 1$ , this does not work well

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– Estimate intermediate quantile  $\hat{Q}_{x}(\tau_0)$ 



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- Estimate intermediate quantile  $\hat{Q}_{x}(\tau_0)$
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- Our extremal random forest uses negative GPD log-likelihood as loss  $L(\theta, Z) = \ell_{\theta}(Z)$  and solves

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(\hat{\sigma}(\mathbf{x}), \hat{\gamma}(\mathbf{x})) = \argmin_{\theta = (\sigma, \gamma)} \sum_{i=1}^{n} w_n(\mathbf{x}, \mathbf{X}_i) \ell_{\theta}(Z_i)
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– Under some assumptions, ERF estimates are consistent

$$
\hat{\theta}(\mathbf{x}) \stackrel{\mathbb{P}}{\rightarrow} \theta(\mathbf{x}), \quad \text{for all } \mathbf{x} \in [-1,1]^p
$$



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#### Simulation study

– Sample  $n = 2000$  iid copies of  $(X, Y)$  from

$$
\begin{cases}\n\mathbf{X} \sim U\left([-1,1]^p\right),\\ \n\left(Y \mid \mathbf{X} = \mathbf{x}\right) \sim s(\mathbf{x})\mathcal{T}_4,\n\end{cases}
$$

where  $s(x) = 1 + 1\{x_1 > 0\}$  and  $\gamma(x) = 1/4$ .

- Compare ERF and GBEX with quantile regression and other extreme value methods
- $-$  On a test data set  $\{{\bf x}_i\}_{i=1}^{n'}$ , evaluate the integrated squared error (ISE)

$$
\mathsf{ISE} = \frac{1}{n'} \sum_{i=1}^{n'} \left( \hat{Q}_{x_i}(\tau) - Q_{x_i}(\tau) \right)^2.
$$



#### [9] N. Gnecco, E. M. Terefe, and S. Engelke. "Extremal random forests". In: J. Amer. Statist. Assoc. (2023). conditionally accepted.

[10] J. Velthoen, C. Dombry, J.-J. Cai, and S. Engelke. "Gradient boosting for extreme quantile regression". In: Extremes (2023). to appear.

#### Simulation study results



#### Extreme quantile regression neural networks (EQRN)

- $-$  If there is sequential dependence as in time series, then this structure can be used in recurrent neural networks
- Predict quantiles of  $Y_t$  (discharge at time t) using past observations

$$
\mathbf{X} = (Y_{t-1}, Y_{t-2}, \dots, X_{t-1}^1, X_{t-2}^1, \dots)
$$

from response on other covariates  $X^1,X^2,\ldots$  (e.g., precipitation at locations 1, 2, etc.)



<sup>[11]</sup> O. C. Pasche and S. Engelke. Neural networks for extreme quantile regression with an application to forecasting of flood risk. Available on https://arxiv.org/abs/2208.07590. 2022.

#### Results for the 2005 flood in Bern



 $-$  Top: One-day-ahead forecasted conditional 100-year return level  $\hat{Q}_{\mathsf{x}}^{100}$  (blue line)

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- $-$  Top: One-day-ahead forecasted conditional 100-year return level  $\hat{Q}_{\mathsf{x}}^{100}$  (blue line)
- Bottom: Ratio of conditional exceedance probability compared to unconditional estimate

$$
\frac{\hat{\mathbb{P}}(\boldsymbol{\Upsilon}>\hat{Q}^{100}\mid\boldsymbol{X}=\boldsymbol{x})}{\hat{\mathbb{P}}(\boldsymbol{\Upsilon}>\hat{Q}^{100})}
$$

## evtGAN: combining EVT and GANs

#### Climate model data

- We use 2000 years of large ensemble simulations with the  $EC$ -Earth global climate model  $[12]$ .
- Present-day climate conditions, stationary in time; see [13] for details.
- Data at  $d = 18 \times 22$  grid points over western Europe.
- We consider annual maxima of precipitation and temperature, giving us  $n = 2000$  observations of a random vector  $\mathbf{Z} = (Z_1, \ldots, Z_d)$ .

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Goal: Realistic simulations of Z for stress testing and extreme event simulation. [14] Here we use  $n_{\text{train}} = 50$  years for training, and  $n_{\text{test}} = 1950$  for evaluation.

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#### Marginal GEV parameters



Figure: GEV parameters for temperature (a-c) and precipitation (d-f): mean parameter  $\mu$  (left), scale parameter σ (center) and shape parameter  $ξ$  (right).

#### Two approaches

#### Classical EVT approach

- $-$  Spatial extreme value theory provides statistical models for  $Z$ , e.g., a spatial max-stable model [15].
- If the region is very large and heterogeneous, such models may not be flexible enough, because of:
	- spatial non-stationarities;
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	- etc.

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#### Our ML approach

- Generative Adversarial Networks (GANs) [16] are a flexible way of learning and sampling from a multivariate distribution Z.
- $-$  They are usually used to sample from image data using convolutional neural networks.
- $-$  We can treat our spatial climatological data  $Z$  as an image.

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#### GAN architecture: generator and discriminator



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#### GANs for extremes

- GANs are trained on the bulk of the distribution.
- $-$  There are two main challenges concerning extremes:
	- accurate extrapolation of the marginal distributions;
	- accurate modeling of the extremal dependence structure.

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Our  $evtGAN$  is copula approach where marginals use  $EVT$  approximations and dependence the structure is generated by the GAN.

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**Input:** Annual maxima  $\mathbf{Z}_i = (Z_{i1}, \ldots, Z_{id})$ ,  $i = 1, \ldots, n$ .

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1. Fit a GEV distribution  $\widehat{G}_j$  to the *j*th marginal with parameters  $(\hat{\mu}_i, \hat{\sigma}_i, \hat{\xi}_i)$ .

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- 2. Normalize empirically to a std uniform distribution to obtain pseudo observations

$$
\mathbf{U}_i=(\widehat{F}_1(Z_{i1}),\ldots,\widehat{F}_d(Z_{id})),\quad i=1,\ldots,n,
$$

where  $\widehat{F}_i$  is the empirical distribution function of the  $Z_{1j}, \ldots, Z_{nj}$ .

3. Train a GAN G on the normalized data  $U_1, \ldots, U_n$ .

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- 4. Generate  $n^*$  new data points  $\mathbf{U}_1^*,\ldots,\mathbf{U}_{n^*}^*$  from G with uniform margins.
- 5. Normalize back to the scale of the original observations

$$
\mathbf{Z}_{i}^{*}=(\widehat{G}_{1}^{-1}(U_{i1}^{*}),\ldots,\widehat{G}_{d}^{-1}(U_{id}^{*})), i=1,\ldots,n^{*}.
$$

**Output:** Set of new generated observations  $\mathbf{Z}_{i}^{*} = (Z_{i1}, \ldots, Z_{id}^{*}), i = 1, \ldots, n^{*}.$ 

#### Bivariate samples of temperature



#### Bivariate samples of precipitation



#### Extremal correlation plots



# Thank you!