# Statistical Distributions and what you can do with them INFN Statistics School, Ischia, 

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## Contents

- General properties
- The main distributions: Poisson and Gaussian
- A quick look at lot of others with one vital fact per distribution
- From Random Numbers to Random distributions


## Some General Properties

Random variable: integer (usually called $r$ ) or real (usually called $x$ ) $P_{r}$ is probability of $r$. Dimensionless numbers. $\sum P_{r}=1$ $P(x)$ is probability density for $x .[P(x)]=[x]^{-1} \cdot \int P(x) d x=1$ Expectation values $<f>=\sum f(r) P_{r}$ or $\int f(x) P(x) d x$

## Measures of Location

Mean: $\mu=<x>$
Mode: $P($ mode $)=\max (P(x))$
Median: $\int{ }^{\text {median }} P(x) d x=0.5$
Measures of Scale
$\sigma=\sqrt{\left.<(x-\mu)^{2}\right\rangle}=\sqrt{<x^{2}>-\langle x\rangle^{2}}$


FWHM=Full Width Half Max
Inter-quartile range

## Other stuff

Skew: $\gamma=\frac{\left\langle(x-\mu)^{3}\right\rangle}{\sigma^{3}}$
Kurtosis: $K=\frac{\left\langle(x-\mu)^{4}\right\rangle}{\sigma^{4}}-3$


Moments: $M_{N}=<x^{N}>, \mu_{N}=<(x-\mu)^{N}>$

## More than one variable: Joint distributions



Two variables
Covariance $\operatorname{Cov}(x, y)=\langle x y\rangle-\langle x\rangle\langle y\rangle$
Correlation $\rho=\frac{\operatorname{Cov}(x, y)}{\sigma_{x} \sigma_{y}}$

## Several variables

Covariance $C_{i j}=\left\langle x_{i} x_{j}>-<x_{i}><x_{j}>\right.$
Correlation $\rho_{i j}=\frac{C_{i j}}{\sigma_{i} \sigma_{j}}$
Diagonals: $C_{i i}=\sigma_{i}^{2}, \rho_{i i}=1$
Can be shown that: $|\rho| \leq 1$

## The Poisson

Memoryless random source. Mean number $\mu$. Actual number $r$

$$
P(r ; \mu)=e^{-\mu} \frac{\mu^{r}}{r!}
$$

Classic example: Geiger counter clicks Also: Prussian soldiers killed by horses. Photomultipliers. Rare decays Counterexamples: photons from lasers. Traffic (especially buses).

Vital fact: $\sigma=\sqrt{\mu}$



Small $\mu$ : 0 is mode $\mu>1$ : peak develops
Distribution has positive skew tail to high values



Large $\mu$ : shape becomes Gaussian

## The Gaussian

$$
\begin{aligned}
& P(x ; \mu, \sigma)=\frac{1}{\sqrt{2 \pi} \sigma} e^{-\frac{(x-\mu)^{2}}{2 \sigma^{2}}} \\
& \text { (Inaccurately) called the 'Bell curve' }
\end{aligned}
$$

$\mu$ is mean and mode and median $\sigma$ is standard deviation $68.27 \%$ of area within $1 \sigma$ so $1 / 3$ of error bars should miss! $95.45 \%$ of area within $2 \sigma$ $99.73 \%$ of area within $3 \sigma$

Gaussian

Describes: large $\mu$ Poisson, measurement errors, height, IQ...
Vital fact: Thanks to Central Limit Theorem: convolution of N random variables $P(x)$ tends to Gaussian for large $N$, irrespective of $P(x)$.

## Demonstrating the CLT

Exercise: using your favourite package (ROOT, Python, Matlab, R, whatever) generate many uniform random numbers and histogram them. Get flat plot, very non-Gaussian. Then generate pairs and add them - get triangular shape. Then triples. Then tens, Looks pretty Gaussian...


Histogram of $x 1+x 2+x 3$



Histogram of $\times 1+\times 2$
ram of $x 1+x 2+x 3+x 4+x 5+x 6+x 7+x 8$


## Central Limit Theorem: the proof

Optional: skip this slide if you're lazy or stupid...
Show: if you convolute $P(x)$ with itself $N(\rightarrow \infty)$ times you get a Gaussian Given $P(x)$, Fourier Transform is $\tilde{P}(k)=\int P(x) e^{i k x} d x=<e^{i k x}>$ Expand and separate: $\left.\left.1+i k\langle x\rangle+\frac{(i k)^{2}}{2!}<x^{2}\right\rangle+\frac{(i k)^{3}}{3!}<x^{3}\right\rangle \ldots$
Take the logarithm, and use $\ln (1+z)=z-\frac{z^{2}}{2}+\frac{z^{3}}{3} \ldots$
Get series in $k: \ln \tilde{P}(k)=(i k) \kappa_{1}+\frac{(i k)^{2}}{2!} \kappa_{2}+\frac{(i k)^{3}}{3!} \kappa_{3}+\ldots$ where the $\kappa_{r}$ ("semi-invariant cumulants of Thiele") are made of expectation values of $x$ to the $r^{\text {th }}$ power. $\kappa_{1}=\langle x\rangle=\mu, \kappa_{2}=\left\langle x^{2}\right\rangle-\langle x\rangle^{2}=\sigma^{2}$, etc
Semi-invariant? Location only changes $\kappa_{1}$, scaling by factor $\alpha, \kappa_{r} \rightarrow \alpha^{r} \kappa_{r}$ Fact: The FT of a convolution is the product of the individual FTs. So the log of the FT of a convolution is the sum of the logs and $K_{r}=N \kappa_{r}$.
To discuss shape, scale by standard deviation $\sqrt{K_{2}}$ $K_{2}^{\prime}=1, K_{r}^{\prime}=K_{r} /{\sqrt{K_{2}}}^{r}=N \kappa_{r} /\left(N \kappa_{2}\right)^{r / 2}$, vanishes as $N \rightarrow \infty$ for $r>2$.
So in the large N limit all $K_{r}$ with $r \geq 3$ vanish, the log of the FT is quadratic:
the FT itself is the exponential of a quadratic, i.e. a Gaussian.
Transforming, the (back) FT of a Gaussian is also a Gaussian. QED.

## Real world Gaussians(1)

Distribution of heights Nice Gaussian distribution


## Real world Gaussians(2)

Distribution of weights .Not really Gaussian. Definite positive skew.


## Real world Gaussians(3)

Distribution of income (per household, in US. Other examples are similar). Totally non-Gaussian.

Distribution of annual household income in the United States 2010 estimate


[^0]
## The Binomial

Probability of $r$ 'successes' from $n$ trials, each with probability $p$.

$$
P(r ; n, p)=\frac{n!}{r!(n-r)!} p^{r} q^{n-r} \quad \text { with } q \equiv 1-p
$$



$P(r ; 0.6,100)$

$P(r ; 0.01,100)$

$\mu=n p \quad \sigma=\sqrt{n p q}$
Limit: $n$ large, $p$ small, $n p=\mu$ fixed $P(r) \rightarrow$ Poisson
Vital Fact: Basically just like tossing coins

## The Uniform

## or Top Hat



Generally, $P(x)=\frac{1}{a}$ between $\mu-a / 2$ and $\mu+a / 2$
Vital fact: Standard Deviation $\sigma=\frac{a}{\sqrt{12}}$

## The Breit-Wigner or Cauchy or Lorentzian

$$
\begin{aligned}
& P(x)=\frac{1}{\pi} \frac{1}{1+x^{2}} \\
& P(E ; M, \Gamma)=\frac{1}{2 \pi} \frac{\Gamma}{(E-M)^{2}+(\Gamma / 2)^{2}}
\end{aligned}
$$

Does not have a standard deviation! integral diverges

FWHM = $\Gamma$ for a Gaussian, FWHM $=2.35 \sigma$, hence use of ' $\sigma$ ' $=\Gamma / 2.35$

Vital fact: Useful for describing measurements that should be Gaussian but aren't

## The log-normal

$P(x)=\frac{1}{\sqrt{2 \pi} \sigma x} e^{-\frac{(\ln x-\mu)^{2}}{2 \sigma^{2}}}$
where $\mu$ and $\sigma$ are mean and sd of $\ln x$


Vital fact: Applies(thanks to CLT) when effect of many factors combine multiplicatively
Example: energy measured in calorimeter with $x=E_{0} \sigma E$

## The Negative Binomial

Coin-tossing again. This time ask 'How many successes before $k$ failures?'



$$
P(r ; k, p)=\frac{(k+r-1)!}{r!(k-1)!} p^{r} q^{k}
$$

Can write factor as $(-1)^{r}{ }_{r} C_{-k}$
(hence unhelpful name)

or as $\frac{\Gamma(k+r)}{\Gamma(k) r!}$
generalise to non-integer $k$
don't ask what that means
All plots here have $p=0.5$
Vital fact: Used to describe events where $\sigma>\sqrt{N}$
i.e. more spread out than Poisson.

## The Weibull

$P(x ; \alpha, \beta)=\alpha \beta(\alpha x)^{\beta-1} e^{-(\alpha x)^{\beta}}$
Devised to describe the lifetime of lightbulbs


'Failure rate' $\propto x^{\beta-1}$

- $\beta<1$ : weak die early ('burn in')
- $\beta=1$ : constant rate (rad. decay)
- $\beta>1$ : aging process
$\alpha$ is just a scale factor



Vital fact: Handy as a way of parametrising rise-and-fall shapes

## The $\chi^{2}$

Much more on this in later lectures!

$$
\chi^{2}=\sum_{1}^{N}\left(\frac{x_{i}-\mu_{i}}{\sigma_{i}}\right)^{2}\left(x_{i} \text { Gaussian }\right)
$$

Measures agreement between $x_{i}$ and $\mu_{i}$
Vital fact: $\overline{\chi^{2}}=N$, but big spread


## Generating Random Distributions

## Starting with the Uniform

Need (pseudo) random number generator for simulations: from Geant4 to Toy Monte Carlo.
All systems seem to contain a function that produces uniform random numbers between 0 and 1 - may be called ran(), $\operatorname{ranf}(), r n d(), r u n i f(), ?, T R a n d o m, T R a n d o m 3 . .$.

100 random numbers


Doesn't look random, does it?
Very easy to see structure!
(Hence the need for Blind Analyses)
Try it yourself!
Extension to other uniform distributions is trivial

## Technical detail

Such functions all based on generator of random integers, then mapped into $[0,1]$.
Classical Method: Linear Congruential Generator (TRandom)
$R_{n+1}=\left(a R_{n}+b\right) \mid c$
with $a, b, c$ suitably chosen. ( $c$ generally $2^{64}$ or $2^{32}$ )
Start with some 'seed' $R_{0}$
(If you want a really random number, use the clock as the seed.)
Drawbacks: repeats with cycle of $2^{64}$ or $2^{32}$ - large but not always large enough. Particular $R$ will never recur till the cycle repeats.
Modern methods: Mersenne Twister(TRandom3) (and its successors).
Large random state from which 62 or 32 bit number extracted.
Even more complicated random numbers used and needed for encryption.

## Other distributions

Suppose you've got a $[0,1]$ random number $U$
For random direction:
$\phi=2 \pi U_{1} \quad \theta=\operatorname{acos}\left(2 U_{2}-1\right)$
The Exponential
Needed to generate decays (with time) and interactions (with distance). If the rate is $r$ then $x=-r \ln (U)$
The Gaussian To get a 'unit Gaussian' $(\mu=0, \sigma=1)$
Lazy way: Add 12 instances of $U$ and subtract 6
Why does this work?
Smart way: Generate $U_{1}$ and $U_{2}$. Form $R=-\ln U_{1}, \theta=2 \pi U_{2}$ Then $R \cos \theta$ and $R \sin \theta$ are both Gaussian random numbers (and uncorrelated!)
Best way: Use the Gaussian generator provided by the system.

## General functions

## 1: Inversion

From desired $P(x)$, form cumulative distribution $C(x)=\int^{x} P\left(x^{\prime}\right) d x^{\prime}$ Generate uniform $u$ in $[0,1]$ and find $x$ such that $C(x)=u$



Example: To generate pdf $P(x)=0.4+0.1 x$ with $x \epsilon[0,2]$
$C(x)=0.4 x+0.05 x^{2}$
$.05 x^{2}+0.4 x-u=0$
$x=\frac{-.4+\sqrt{.16+.2 u}}{0.1}$
Works great if you can (1) integrate $P(x)$ to get $C(x)$ and (2) invert $u=C(x)$ to get $x=C^{-1}(u)$
If not possible analytically, numerical methods may be used

## General functions

Generate $x$ uniformly over range
Generate $r$ uniformly between 0 and $M$, where $M \geq \max (P(x))$
If $P(x)>r$, accept.
Else reject and try again


Works easily for multidimensional functions.
If $M$ is overestimated, method is still valid (just a hit in the efficiency)
Can be very inefficient if $P(x)$ has sharp peaks-
may be improved by generating $x$ according to some $P_{0}$ and using $P(x) / P_{0}(x)$ in the acceptance comparison

## General functions

3: Weighting

Not all events need to be equal!
Generate $x$ uniformly and weight the event by $P(x)$
when filling histograms, forming sums, etc, include the weight.
Can be effective when simulating low-probability processes that reject a lot of events.
More work, but not as hard as it looks.

Doesn't always help...
Poisson error on a weighted number is $\sqrt{N \overline{w^{2}}}$, always bigger than $\sqrt{N} \bar{w}$, i.e. error worse than pure Poisson $\sigma=\sqrt{N}$.

If weights all much the same, not a problem.
If a few events with enormous weights dominate, get big statistical errors.

## Conclusions

There are many distribution for you to use - very big toolkit
Sometimes founded on dynamics of the problem
Sometimes empirical, found by experience to have useful behaviour in particular circumstances

Be open-minded and on the lookout for new ones!


[^0]:    Source U.S. Census Burean, Current Pcpulation Survey, 2011 Annual Social and Econcemic Supplement

